

1 The Coreset Variational Bayes (CVB) Algorithm for  
2 Mixture Analysis

3 Qianying Liu<sup>a</sup>, Clare A McGrory<sup>a</sup>, Peter WJ Baxter<sup>a,b</sup>

4 <sup>a</sup>*Centre for Applications in Natural Resource Mathematics*  
5 *School of Mathematics, The University of Queensland*  
6 *Brisbane, Australia, Qld 4072.*  
7 *Tel.: +61 7 336 51385*  
8 *Fax: +61 7 336 51477*  
9 <sup>b</sup>*Queensland University of Technology*  
10 *Brisbane, Australia, Qld 4001.*

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11 **Abstract**

The pressing need for improved methods for analysing and coping with big data has opened up a new area of research for statisticians. Image analysis is an area where there is typically a very large number of data points to be processed per image, and often multiple images are captured over time. These issues make it challenging to design methodology that is reliable and yet still efficient enough to be of practical use. One promising emerging approach for this problem is to reduce the amount of data that actually has to be processed by extracting what we call coresets from the full dataset; analysis is then based on the coreset rather than the whole dataset. Coresets are representative subsamples of data that are carefully selected via an adaptive sampling approach. We propose a new approach called coreset variational Bayes (CVB) for mixture modelling; this is an algorithm which can perform a variational Bayes analysis of a dataset based on just an extracted coreset of the data. We apply our algorithm to weed image analysis.

12 *Keywords:* Mixture modelling, Coresets, Variational Bayes, Image  
13 Analysis, Bayesian Statistics

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14 **1. Introduction**

15 Finite mixture models are applied in diverse areas of science and provide a straightforward, but flexible extension of classical parametric models (Fruhworth-Schnatter, 2006). They are used in situations where it is

18 thought that there is more than one sub-population giving rise to points in  
19 the dataset. The sub-populations are each represented by one of the com-  
20 ponents that comprise the mixture. An obvious scenario where this type  
21 of model would be suitable is image analysis: we can associate ranges of  
22 intensity level present in the image with components of the mixture.

23 We take a Bayesian approach to our inference. In Bayesian analysis the  
24 posterior distribution of the unknown parameters can be very difficult to esti-  
25 mate. The most commonly used Bayesian method for estimating the param-  
26 eters is Markov chain Monte Carlo (MCMC), e.g. a Gibbs sampling approach  
27 (Alston et al., 2012). However, while MCMC-based approaches are very ac-  
28 curate, the computational requirements associated with this approach can  
29 be prohibitive for very large datasets. Less computationally demanding al-  
30 gorithms for fitting the mixture models are alternative approximate Bayesian  
31 inference techniques such as variational Bayes (VB) (McGrory & Tittering-  
32 ton, 2007; Ormerod & Wand, 2010) and approximate Bayesian computation  
33 (ABC) (Marin et al., 2012).

34 Even though the standard VB method is highly computationally efficient  
35 in comparison to MCMC, in some cases it still might be too time-consuming  
36 for very large data problems if analysis is required within short time-frames.  
37 Consider for instance the weed-crop imaging application that we will explore  
38 in this article. Images have an extremely large number of pixels, there will  
39 likely be multiple images to analyse, and estimates are needed reasonably  
40 quickly. In order to achieve this, time-efficient techniques are required. An-  
41 other avenue is to reduce the volume of data that actually has to be processed  
42 in the first place. Removing part of the dataset before running the analysis is  
43 a less drastic thing to do than we might think when we consider that much  
44 of the dataset gives us the same or very similar information. For example,  
45 consider trying to fit a mixture model to an image, we do not necessarily  
46 have to analyse all of the observed intensity levels present in the full dataset  
47 to obtain a good estimate of the mean for that particular cluster. The coresets  
48 approach (Feldman et al., 2011) is a data-reduction algorithm. It involves  
49 extracting a representative subsample of the dataset which can then be anal-  
50 ysed in order to make inference about the whole dataset. In Feldman et al.  
51 (2011) the coresets approach was used within a classical mixture modelling  
52 framework with good results. In Ahfock et al. (2014) the Gibbs sampler  
53 was modified for use with coresets; the resulting algorithm was called the  
54 weighted Gibbs sampler. This was shown to give very good results when  
55 applied to analysing satellite image data and it drastically reduced the com-

56 putation time required for the analysis. In a similar spirit, we wish to modify  
57 the VB algorithm to make it suitable for use with coresets of data. Since  
58 VB is more time-efficient than the Gibbs sampler, combining the concept  
59 of coresets with VB will result in even greater reductions in computational  
60 burden. We refer to this new algorithm as coreset variational Bayes (CVB).

61 It could be argued that a spatial mixture model is more appropriate  
62 than a finite mixture model for modelling a dataset such as an image which  
63 contains spatial information (see e.g. McGrory et al. (2012)) as this might  
64 slightly increase the clustering accuracy. The reason we do not incorporate a  
65 spatial component into our modelling is that we cannot afford the huge extra  
66 computational burden this would incur since we require a very time efficient  
67 approach for big data settings. In this way there is a trade-off between these  
68 two aspects.

69 Weeds are defined as being plants which have originated in and continue  
70 to evolve in a natural environment; but they are problematic because they  
71 do so in a manner which interferes with the growth of crops or other agricul-  
72 ture related activities (Zimdahl, 2009). Weeds are generally better able to  
73 compete than crops for resources like minerals, light and water. This leads  
74 to a lot of waste of agricultural investment in cases where weed plants are  
75 present, because most of these valuable inputs would be used up by them  
76 instead of by the valuable crops. There are many other harmful problems as-  
77 sociated with weeds, such as the harbouring of extra pests in the area which  
78 can cause plant diseases that may infect the crops in the region. Hence, it is  
79 important to be able to effectively manage weeds in farming regions if a na-  
80 tion's agriculture industry is to remain competitive in international markets  
81 (Sindenab et al., 2004).

82 A growing area of research that has the potential to be very useful in weed  
83 management is the use of statistical methodology to analyse images of weeds  
84 used in agriculture trials. This might be of use in projects where the aim  
85 is to assess and compare the effectiveness of different chemical weed killers  
86 for instance. If calculation of the proportion of a plot of land that is weed,  
87 soil or plant after chemical applications can be done by analysing an image,  
88 this will save time by removing the need for researchers to go out into the  
89 fields and count the live and dead plants by hand. Due to the large number  
90 of pixels present in a typical image, and the fact that there will most likely  
91 be multiple images to analyse in a given trial, it is important to find a very  
92 time efficient algorithm for processing this type of data (see e.g. Kargar and  
93 Shirzadifar (2013)).

94 **2. Finite mixture model**

95 Mixture models provide an excellent and flexible way to represent com-  
 96 plex distributions. The mixture model we fit to our data comprises a linear  
 97 combination of standard mixture models, these are called the components  
 98 of the mode. Each component has a corresponding mixture weight which  
 99 reflects the expected proportion of the data that might be captured by that  
 100 particular component. In our finite mixture model for a set of continuous  
 101 observations  $\mathbf{y}_1, \dots, \mathbf{y}_n$ , assume that the observations are all generated *i.i.d.*  
 102 (independent and identically distributed) from a random variable  $\mathbf{Y}$ , which  
 103 follows a mixture of  $K$  independent Gaussian distributions. In the missing-  
 104 data interpretation of the mixture we introduce an unobserved indicator vari-  
 105 able  $\mathbf{z}_{ij}$  for each observation; this identifies the component allocations of our  
 106 observations by taking value 1 if observation  $i$  is from component  $j$ , and 0  
 107 otherwise. Since these indicators are unknown to us, the model is therefore  
 108 a missing-data model. For each component,  $j$ , the value of  $\rho_j$  is the relevant  
 109 weight for that particular  $j$ th component. The model density is given by:

$$p(\mathbf{y}, \mathbf{z} | \boldsymbol{\theta}) = \prod_{i=1}^n \prod_{j=1}^K \left\{ \rho_j N_d(\mathbf{y}_i; \boldsymbol{\mu}_j, \mathbf{T}_j^{-1}) \right\}^{z_{ij}}.$$

110 Here,  $N_d(\cdot, \cdot)$  represents the  $d$ -dimensional multivariate normal density,  
 111 where  $\boldsymbol{\mu} = (\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_K)$  and  $\mathbf{T} = (\mathbf{T}_1, \dots, \mathbf{T}_K)$ , and  $\mathbf{T}_j$  denotes the  $j$ th  
 112 precision matrix, which is the inverse of the  $j$ th covariance matrix.

113 **3. Bayesian priors**

We follow the standard Bayesian conjugate prior setting (Alston et al.,  
 2012) for this model, and choose hyper-parameters such that they corre-  
 spond to non-informative prior settings, thus allowing information contained  
 in the dataset to have more influence over the fit. The weight coefficients are  
 assigned Dirichlet prior distributions:

$$p(\boldsymbol{\rho}) = Dir(\boldsymbol{\rho}; \alpha_1^{(0)}, \dots, \alpha_K^{(0)}).$$

114 The prior distributions of the means conditioned on the covariance matrices  
 115 are independent multivariate normal distributions:

$$p(\boldsymbol{\mu}|\mathbf{T}) = \prod_{j=1}^K N_d \left( \boldsymbol{\mu}_j; \mathbf{m}_j^{(0)}, (\beta_j^{(0)} \mathbf{T}_j)^{-1} \right).$$

The prior of the precision matrices are given by Wishart distributions:

$$p(\mathbf{T}) = \prod_{j=1}^K W \left( \mathbf{T}_j; \mathbf{v}_j^{(0)}, \boldsymbol{\Sigma}_j^{(0)} \right).$$

Therefore, the joint distribution would finally be:

$$p(\mathbf{y}, \mathbf{z}, \boldsymbol{\theta}) = p(\mathbf{y}, \mathbf{z} | \boldsymbol{\theta}) p(\boldsymbol{\rho}) p(\boldsymbol{\mu}|\mathbf{T}) p(\mathbf{T}).$$

116 The quantities of  $\{\alpha_j^{(0)}\}$ ,  $\{\mathbf{m}_j^{(0)}\}$ ,  $\{\beta_j^{(0)}\}$  and  $\{\boldsymbol{\Sigma}_j^{(0)}\}$  are all hyper-parameters.  
117

## 118 4. Bayesian posterior distributions

### 119 4.1. The variational approach

120 The variational Bayesian method is a time-efficient approach for estimating  
121 Bayesian mixture models (Faes et al., 2011; McGrory & Titterton,  
122 2007; Wand et al., 2012) and can be thought of as an alternative to MCMC.  
123 The main difference between the two approaches is that instead of estimating  
124 the parameter directly by sampling from the posterior distribution, varia-  
125 tional Bayesian methods approximate it. This is done by artificially "intro-  
126 ducing" a more amenable distribution  $q(\boldsymbol{\theta}, \mathbf{z})$  which is often referred to as  
127 the variational approximating function. This function will end up becoming  
128 an approximate the joint conditional distribution of  $\boldsymbol{\theta}$  and  $\mathbf{z}$  given the ob-  
129 servations  $\mathbf{y}$  after the variational method is applied to it. In the following  
130 we explain how  $q(\boldsymbol{\theta}, \mathbf{z})$  should be chosen and integrated into the variational  
131 framework in order to achieve this outcome.

132 The distribution  $q(\boldsymbol{\theta}, \mathbf{z})$  is chosen to minimise the Kullback-Leibler(KL)  
133 divergence between the approximating density  $q(\boldsymbol{\theta}, \mathbf{z})$  and the true joint den-  
134 sity  $p(\boldsymbol{\theta}, \mathbf{z} | \mathbf{y})$ . In doing so, we are trying to obtain a relatively tight lower  
135 bound on the marginal density,  $p(\mathbf{y})$ . Essentially we manipulate and re-  
136 xpress the joint density to allow us to introduce the variational approximat-  
137 ing function in such a way that we can then use a maximisation approach to  
138 estimate parameters of that target approximating function (see also McGrory

139 & Titterington (2007)). We begin then by showing that the joint density is  
 140 lower bounded as:

$$\begin{aligned}
 \log p(\mathbf{y}) &= \log \int \sum_{\{\mathbf{z}\}} p(\mathbf{y}, \mathbf{z}, \boldsymbol{\theta}) d\boldsymbol{\theta} \\
 &= \log \int \sum_{\{\mathbf{z}\}} q(\boldsymbol{\theta}, \mathbf{z}) \frac{p(\mathbf{y}, \mathbf{z}, \boldsymbol{\theta})}{q(\boldsymbol{\theta}, \mathbf{z})} d\boldsymbol{\theta} \\
 &\geq \int \sum_{\{\mathbf{z}\}} \log \frac{p(\mathbf{y}, \mathbf{z}, \boldsymbol{\theta})}{q(\boldsymbol{\theta}, \mathbf{z})} d\boldsymbol{\theta}. \quad \text{by Jensen's inequality} \quad (1)
 \end{aligned}$$

141 Another way of viewing this is that finding the tightest lower bound is the  
 142 same as minimising the Kullback-Leibler divergence between the variational  
 143 distribution and the true target posterior. It is exactly minimised when we  
 144 take  $q(\boldsymbol{\theta}, \mathbf{z}) = p(\boldsymbol{\theta}, \mathbf{z}|\mathbf{y})$ . However, as we are trying to simplify the problem,  
 145  $q(\boldsymbol{\theta}, \mathbf{z})$  should be a close enough approximation to the true density, yet have  
 146 a simple form for computational purposes. Normally, to achieve this  $q(\boldsymbol{\theta}, \mathbf{z})$   
 147 is restricted to have the factorised form  $q(\boldsymbol{\theta}, \mathbf{z}) = q_{\boldsymbol{\theta}}(\boldsymbol{\theta})q_{\mathbf{z}}(\mathbf{z})$ .

148 Unlike the MCMC approach, the variational method approximates the  
 149 parameters in the finite mixture model. This difference may cause a slight  
 150 decrease in accuracy of the variational method when compared with MCMC.  
 151 It has been demonstrated that in many contexts, including mixture modelling  
 152 (e.g. McGrory & Titterington (2007); Wand et al. (2012); Faes et al. (2011))  
 153 that the variational method can largely reduce operating time and yet the loss  
 154 in accuracy that arises from the approximation is not terribly great. Indeed  
 155 the approximate result is typically adequate for practical purposes. When  
 156 computational efficiency is an important consideration, it is worthwhile for  
 157 us to pursue this approximate approach rather than MCMC.

#### 158 4.2. Variational posterior

159 After we maximise the lower bound (equation (1)), the posteriors are

$$q_{\boldsymbol{\rho}}(\boldsymbol{\rho}) = Dir(\boldsymbol{\rho}; \alpha_1, \dots, \alpha_k),$$

$$q_{\boldsymbol{\mu}|\mathbf{T}}(\boldsymbol{\mu}|\mathbf{T}) = \prod_{j=1}^K N_d(\boldsymbol{\mu}_j; \mathbf{m}_j, (\beta_j \mathbf{T}_j)^{-1}),$$

and

$$q_{\mathbf{T}}(\mathbf{T}) = \prod_{j=1}^K W(\mathbf{T}_j; \mathbf{v}_j, \Sigma_j),$$

160 the hyperparameters will be updated as:

$$\alpha_j = \alpha_j^{(0)} + \sum_{i=1}^n q_{ij} \quad (2)$$

$$\beta_j = \beta_j^{(0)} + \sum_{i=1}^n q_{ij} \quad (3)$$

$$\mathbf{m}_j = \frac{\beta_j^{(0)} \mathbf{m}_j^{(0)} + \sum_{i=1}^n q_{ij} \mathbf{y}_i}{\beta_j} \quad (4)$$

$$\Sigma_j = \Sigma_j^{(0)} + \sum_{i=1}^n q_{ij} \mathbf{y}_i \mathbf{y}_i^T + \beta_j^{(0)} \mathbf{m}_j^{(0)} \mathbf{m}_j^{(0)T} - \beta_j \mathbf{m}_j \mathbf{m}_j^T \quad (5)$$

$$v_j = v_j^{(0)} + \sum_{i=1}^n q_{ij}, \quad (6)$$

161 with  $q_{ij}$  being the variational posterior expected probability that the in-  
162 dicator variable  $z_{ij} = 1$ . The form of  $q_{ij}$  is:

$$\begin{aligned} q_{ij} &= \frac{\exp \left\{ \langle \log \rho_j \rangle + \frac{1}{2} \{ \langle \log |\mathbf{T}_j| \rangle \} \right\} - \frac{1}{2} \text{tr} \left( \langle \mathbf{T}_j \rangle (\mathbf{y}_i - \mathbf{m}_j) (\mathbf{y}_i - \mathbf{m}_j)^T + 1/\beta_j \right) \mathbf{I}_d}{s_i} \\ &= \frac{\phi_{ij}}{s_{ij}}, \end{aligned} \quad (7)$$

163 where  $s_{ij}$  is the normalization constant and  $\langle \cdot \rangle$  denotes the expected val-  
164 ues required to evaluate the expressions in Equation 7. These are given by:

$$\langle \boldsymbol{\mu}_j \rangle = \mathbf{m}_j,$$

$$\langle \mathbf{T}_j \rangle = v_j \Sigma_j^{-1}$$

$$\langle |\log \mathbf{T}_j| \rangle = \sum_{s=1}^d \Psi \left( \frac{v_j + 1 - s}{2} \right) + d \log(2) + \log |\Sigma_j|$$

$$\langle \log(\rho_j) \rangle = \Psi(\hat{\alpha}_j) + \Psi(\hat{\alpha}_\cdot),$$

165 where  $\Psi(\cdot)$  is the digamma function and  $\hat{\alpha}_\cdot = \sum_j \hat{\alpha}_j$ .

166 4.3. *The Standard Variational Bayes (VB) algorithm*

167 In Algorithm 1 we outline the pseudo-code for the VB algorithm for ob-  
168 taining the posterior estimates. We draw the reader’s attention to what we  
169 call the component elimination property of VB. By this we mean that the  
170 algorithm can automatically determine the complexity of the model since:

- 171 1. the initial value for the number of components is set to be greater than  
172 what the user would reasonably expect in the final fit,
- 173 2. components which converge to have similar estimated parameter val-  
174 ues will be dominated by only one of them, then components with  
175 small mixing weights can be removed leading to automatic complexity  
176 assessment.

177 This feature is very useful in applications involving large amounts of data  
178 as it greatly reduces computing costs: no need to perform different runs for  
179 various numbers of mixture components and compare them, which is what  
180 we would have to do if we used an MCMC approach in order to estimate  
181 the dimension. The other alternative would be to use the computationally  
182 burdensome reversible jump MCMC. This feature is also particularly useful  
183 for practical applications where the operator would like the analysis to run  
184 in an unsupervised manor because there is no need for the user to manually  
185 control searches over dimensionality.

186 Of course initial settings for the hyperparameters and epsilon (the thresh-  
187 old for component removal) can have some influence on estimation some  
188 cases, although previous research has shown the method to be generally fairly  
189 robust to initial parameter settings. We chose values for the hyperparame-  
190 ters  $\alpha^{(0)}, \beta^{(0)}, \Sigma^{(0)}, v^{(0)}, \mathbf{m}^{(0)}$  to correspond to vague non-informative priors.  
191 Note that these are the hyperparameters of the Gaussian mixture model, and  
192 they need to be chosen in the initial settings of the variational algorithm. In  
193 this we follow the standard guidelines for prior settings used in any Bayesian  
194 analysis. Naturally, if a user has specific prior knowledge in their particular  
195 application, they might decide an informative prior is suitable in their case,  
196 this is a user driven choice.

197 Note that the choice of epsilon determines how small a component’s al-  
198 located weighting has to be at a given iteration of the algorithm in order for  
199 it to be selected for removal, and again is the user’s choice, within reason  
200 of course. The simplest, and perhaps ”safest” choice is to set that equal  
201 to 1. That is a ”safe” option because clearly once number of allocations is

202 less than 1, the user can be sure that they are not excluding a component  
 203 that has any real significant contribution within the model. However, that  
 204 might not be the most time efficient choice, if you imagine a dataset with  
 205 thousands of observations, it might be that any component with less than  
 206 say 50-100 observations allocated is unlikely to be useful and on the way  
 207 to being eventually removed with a gradual reduction in allocations at each  
 208 subsequent iteration. However, choosing epsilon in the range 50-100 might be  
 209 less "safe" because it does incur higher risk that a component that is in fact  
 210 potentially useful in representing some features of the data will get removed  
 211 in error. Some users might decide on a proportional value of the size of the  
 212 observation set which can be used as the cut off, e.g. 1% of the dataset size.  
 213 Users must select a suitable value according to their particular application.

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**Algorithm 1:** The standard VB algorithm

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Set initial number of components  $K$ .

Set initial values for hyperparameters  $\alpha^{(0)}, \beta^{(0)}, \Sigma^{(0)}, v^{(0)}, \mathbf{m}^{(0)}$ .

Specify initial allocation of observations to components and get  $q_{ij}$ .

**while** Not Converged **do:**

Update variational posterior expressions for model parameters: equations 2-6

Update variational posterior for  $q_{ij}$ : equation 7

**if** any component has a mixing weight  $\leq \varepsilon$

remove the component from model

**end if**

**if** the algorithm has converged

exit loop.

**end if**

**end while**

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215 **5. Adapting the Variational Bayes Approach for use with Coresets**  
216 **of Data**

217 In Ahfock et al. (2014) it was shown how the Gibbs sampler could be  
218 adapted for use with coresets of data. In a similar spirit we will adapt the  
219 VB method for use with coresets. We first describe the basic procedure for  
220 finding coresets, as outlined in Feldman et al. (2011).

221 *5.1. Finding coresets*

222 The coreset method described in Feldman et al. (2011) can be used to find  
223 an appropriate weighted subset to represent the information in the complete  
224 dataset. The starting point is to first sample uniformly a small number of  
225 points, then remove half of the data points which are closest to the sampled  
226 points. Next sample again from the rest of the points and remove half of  
227 the points lying closest until all of the data points are labeled as removed or  
228 sampled.

229 By doing this, we construct a hierarchy of data points and the importance-  
230 weight of the sampled points is associated with the log-likelihood. The  
231 weights are set to be optimal if the estimated log-likelihood is of the least  
232 variance. This construction of a sampled set gives a higher probability to  
233 observations that are further away from the initial cluster center, and the  
234 sampling bias would be fixed by adapting the weight which is to be associated  
235 with the sampling probabilities. We can then finally build an appropriate  
236 coreset from the whole dataset based on the weights.

237 This algorithm for coreset construction is more formally summarised in  
238 pseudo-code form in Algorithm 2.

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**Algorithm 2:** Algorithm for finding a coreset (see Feldman et al. (2011))

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**input:** Whole dataset  $\mathbf{D}$ ,  $\epsilon$ ,  $\delta$ ,  $K$ .

**set:**  $\mathbf{D}' \leftarrow \mathbf{D}$ ;  $\mathbf{B} \leftarrow \emptyset$

Specify initial allocation of observations to components.

**while**  $|\mathbf{D}'| \geq 10dK \ln(1/\delta)\epsilon$  **do** :

    Sample set  $\mathbf{S}$  of  $\beta = 10dK \ln(1/\delta)$  points uniformly at random from  $\mathbf{D}'$ ;

    Remove  $\lceil |\mathbf{D}'|/2 \rceil$  points  $\mathbf{x} \in \mathbf{D}'$  closest to  $\mathbf{S}$  (i.e., minimising  $\text{dist}(\mathbf{x}, \mathbf{S})$ ) from  $\mathbf{D}'$ ;

    Set  $\mathbf{B} \leftarrow \mathbf{B} \cup \mathbf{S}$ ;

Set  $\mathbf{B} \leftarrow \mathbf{B} \cup \mathbf{D}'$

239

**for** every  $b \in \mathbf{B}$  **do**

$\mathbf{D}_b \leftarrow$  the points in  $\mathbf{D}$  whose closest point  $\mathbf{B}$  is  $b$ .

**for** every  $b \in \mathbf{B}$  and every  $\mathbf{x} \in \mathbf{D}_b$  **do**

$$m(\mathbf{x}) \leftarrow \lceil \frac{5}{|\mathbf{D}_b|} + \frac{\text{dist}(\mathbf{x}, \mathbf{B})^2}{\sum_{\mathbf{x}' \in \mathbf{D}} \text{dist}(\mathbf{x}, \mathbf{B})^2} \rceil;$$

Pick a non-uniform random sample  $\mathbf{C}$  of  $10 \lceil dk |\mathbf{B}^2 \ln(1/\delta) / \epsilon^2 \rceil$  points from  $\mathbf{D}$ ,

where for every  $\mathbf{x}' \in \mathbf{C}$  and  $\mathbf{x}' \in \mathbf{D}$ , we have  $\mathbf{x}' = \mathbf{x}$  with probability

$$m(\mathbf{x}) / \sum_{\mathbf{x}' \in \mathbf{D}} m(\mathbf{x}');$$

**for** each  $\mathbf{x}' \in \mathbf{C}$  **do**  $\gamma(\mathbf{x}') \leftarrow \frac{\sum_{\mathbf{x} \in \mathbf{D}} m(\mathbf{x})}{|\mathbf{C}| \cdot m(\mathbf{x}'')}$

**output:** Coreset  $\mathbf{C} = \{(\gamma(\mathbf{x}_1), \mathbf{x}_1), (\gamma(\mathbf{x}_2), \mathbf{x}_2), \dots, (\gamma(\mathbf{x}_{|\mathbf{C}|}), \mathbf{x}_{|\mathbf{C}|})\}$ .

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240 5.2. VB inference using coreset sampling

241 In this section we propose a new algorithm in which we adapt the varia-  
 242 tional Bayes method in order that it can be used in conjunction with a coreset  
 243 sampling approach. We will use the standard prior settings as described in

244 section 3 and the posterior of the model will be adjusted using the coresets  
 245 samples and coresets weights. This novel modification of the algorithm makes  
 246 it suitable for use in analysing a coresets of the image. The update equations  
 247 that have to be adapted for use with the coresets data are equations 2-7. In  
 248 those equations  $\mathbf{y}_i$  is replaced by  $\hat{\mathbf{y}}_i$ , where  $\hat{\mathbf{y}}_i$  corresponds to the weighted  
 249 observations and is defined as follows:

$$\hat{\mathbf{y}}_i = \frac{\gamma_i \times \mathbf{y}_i}{\frac{1}{n} \sum_{i=1}^n \gamma_i}.$$

250 The expected values required to update the expressions remain unaltered  
 251 from the form they take in the standard VB algorithm. Pseudo-code for  
 252 the weighted VB algorithm which we call coresets variational Bayes (CVB) is  
 253 outlined in Algorithm 3.

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**Algorithm 3:** Coresets Variational Bayes (CVB) Algorithm

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**input:**  $C = \{(\gamma(\mathbf{x}_1), \mathbf{x}_1), (\gamma(\mathbf{x}_2), \mathbf{x}_2), \dots, (\gamma(\mathbf{x}_N), \mathbf{x}_N)\}$  from Algorithm 2.  
 Set initial values for hyper-parameters,  $\varepsilon, \alpha^{(0)}, \beta^{(0)}, \Sigma^{(0)}, v^{(0)}, m^{(0)}$ .  
 Specify initial allocation of observations to components via initial  $q_{ij}$ .

**while** Not Converged, **do**:  
 Update modified variational posterior expressions for model parameters  
 Update modified variational posterior for  $q_{ij}$   
 254 **if** any component has a mixing weight  $\leq \varepsilon$   
 remove the component from model  
**end if**  
**if** the algorithm has converged  
 exit loop.  
**end if**  
**end while**

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255 **6. Application to Weed-Crop Image Segmentation Using Coreset**  
256 **Variational Bayes (CVB)**

257 *6.1. Weed Image*

258 We illustrate the effectiveness of the CVB algorithm for analysing an  
259 image of a weed plant amongst soil and dead leaves, see Fig 1. The size of the  
260 coreset is 2599 pixels, which is 1/100 of the size of the original dataset. The  
261 idea is to use the CVB algorithm to segment the image and classify the pixels  
262 as representing either living plant, background soil, or dead leaves. This is  
263 the type of classification that would be required for research purposes in  
264 agricultural trials. In running the analysis, the hyper-parameters are chosen  
265 to correspond to vague prior settings.

266 Figure 2 shows the segmentation of each component into the three differ-  
267 ent types of matter. As we can see, the area of the living part of the weed is  
268 clearly defined, and the algorithm can also distinguish between soil and dead  
269 leaves. The numerical segmentation results are in Table 1 and for comparison  
270 we also show the results that we would have obtained from an analysis of the  
271 full dataset. There is close agreement between these results showing that the  
272 coreset modification of the VB algorithm is reliable and useful. Significantly,  
273 the VB coreset algorithm is greatly more time-efficient than the standard VB  
274 as it runs around 18 times faster (around 40 minutes for CVB compared to  
275 around 12 hours for VB on the full dataset). This is very impressive when  
276 we consider how similar the final results were. Hence CVB is a practical and  
277 useful alternative to standard VB for the image segmentation based on finite  
278 Gaussian mixture models.



Figure 1: The original image of a weed against background soil to be analysed.

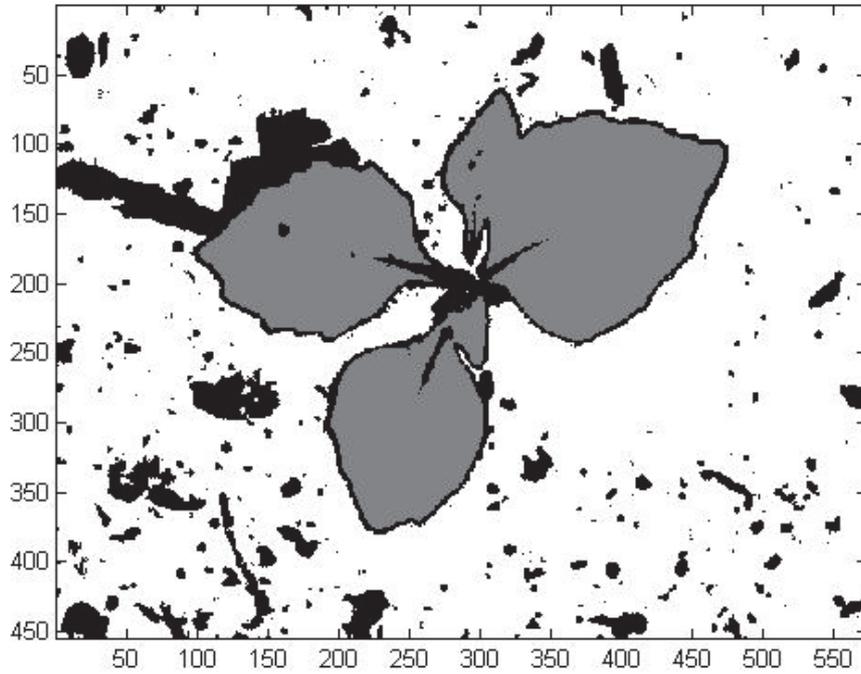


Figure 2: Result of the CVB analysis of the weed image. The pixels in the image have been segmented into 3 different components representing the types background soil, plant and dead leaf.

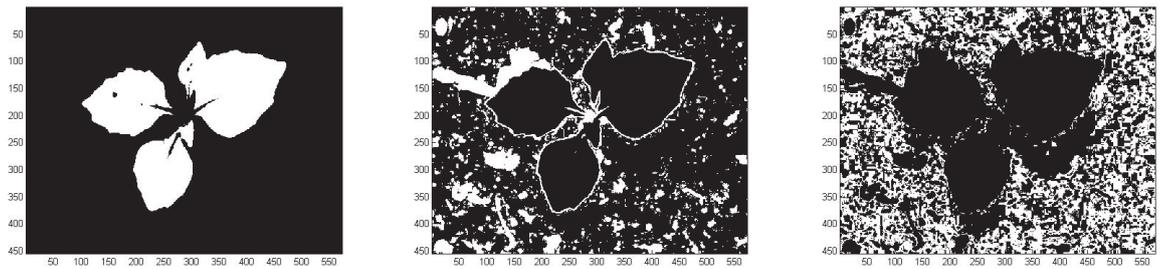


Figure 3: The three different components identified via the CVB algorithm shown in three separate plots. The first plot shows, in white, the area of living plant (weed), the second shows dead leaves. The third plot shows regions of background soil.

Table 1: Result comparison

Component type	standard VB		VB coresets	
	mixing weight	mean	mixing weight	mean
background soil	0.7005	$\begin{bmatrix} 0.1075 \\ 0.0986 \\ 0.0941 \end{bmatrix}$	0.6989	$\begin{bmatrix} 0.1093 \\ 0.0999 \\ 0.0952 \end{bmatrix}$
plant	0.1765	$\begin{bmatrix} 0.3200 \\ 0.4396 \\ 0.1753 \end{bmatrix}$	0.1894	$\begin{bmatrix} 0.3233 \\ 0.4422 \\ 0.1810 \end{bmatrix}$
dead leaves	0.1231	$\begin{bmatrix} 0.2660 \\ 0.2447 \\ 0.1986 \end{bmatrix}$	0.1117	$\begin{bmatrix} 0.2768 \\ 0.2561 \\ 0.2059 \end{bmatrix}$

## 279 7. Discussion

280 We have presented a new algorithm (CVB) for analysing data using vari-  
281 ational Bayes based on a representative coresets of the data. This allows us  
282 to perform reliable inference in a highly time-efficient way. However, the  
283 running time after algorithm modification is still around 40 minutes for a  
284 medium size (67KB) image. While this is good in comparison to other ex-  
285 isting approaches, there is still much scope for further research if these ideas  
286 are to be put to routine use. Consider that agricultural activities may need  
287 weed detection means applied on a large area of land, and thus the image  
288 can be more than 1GB, and contain even more detail than the presented ex-  
289 ample. One option to explore might be looking at improving computational  
290 speed through the use of more efficient programming algorithms, more so-  
291 phisticated computers or GPU programming for increased efficiency Suchard  
292 et al. (2010). This is a topic for future research.

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